

# Marginal versus conditional effects: does it make a difference?

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# Overview

- In observational and experimental studies, the goal may be to estimate the effect of a binary treatment.
- Researchers often adjust for baseline covariates using logistic regressions. The coefficient of treatment in this model is interpreted as the « effect » (in terms of the log-odds ratio).
- The parameters in a logistic regression are uncollapsible – this means that if you add a covariate that is related to the treatment and/or outcome, the true value of the coefficient of treatment changes.
- My view (which is pioneered by many working in both clinical trials and causal inference) is that it is preferable to try and estimate the marginal odds ratio.

# Overview

- In these slides, I present a simple randomized setting where we compare the difference between the true conditional odds ratio and the true marginal odds ratio.
- We see that these parameters (compared on a relative scale) can differ substantially but that this will depend on the **associations** between the variables and the outcome and on the **frequency** of the outcome.

# Setting

- Suppose we have baseline data,  $C$ , for every subject.
- Binary exposure indicator,  $A$ .
- Outcome  $Y$ .
- Interest lies in estimating the effect of exposure  $A$  on outcome  $Y$ .

# Traditional effect measures

- In traditional statistical approaches, we propose a model that represents the outcome process, i.e.  $E(Y|A, C)$ .
  - E.g. A linear/logistic regression
- This model is made conditional on treatment type and all covariates deemed necessary to unconfound the effect estimation (or improve efficiency in a randomized trial).
- Finally, the coefficient of the treatment covariate is interpreted as the effect of interest.

# Marginal effect measures

- Each subject has the potential to be exposed to either intervention level (*positivity*).
- Therefore, each subject has two potential outcomes (or *counterfactuals*)  $Y^1$  and  $Y^0$ .
  - We only observe the one that corresponds with the intervention actually received.

# Adjusted vs conditional

- Marginal effects defined using potential outcomes have a **causal interpretation** regardless of the estimation methods used.
- The marginal odds ratio,  $\frac{P(Y^1=1)/P(Y^1=0)}{P(Y^0=1)/P(Y^0=0)}$ , contrasts the odds of obtaining the outcome had everyone in the population received intervention 1 versus 0.
- To estimate marginal effects, it might still be necessary to **adjust for confounders**.
- Note that there are many available methods to estimate the marginal odds ratio while adjusting for confounders (e.g. inverse propensity score weighting, G-Computation, and Targeted Maximum Likelihood Estimation).

# Problems with traditional methods

## Non-collapsibility

- If you use a logistic regression model to estimate an effect on a binary outcome, the resulting odds ratio is *conditional* on the covariates in the model.
- This is different from the *marginal* odds ratio,  $\frac{P(Y^1=1)/P(Y^1=0)}{P(Y^0=1)/P(Y^0=0)}$ .



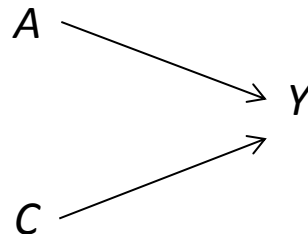
# Problems with traditional methods

## Why is this a problem?

- Suppose you have a **randomized controlled trial** with two arms and a binary outcome. (Full adherence, no dropout, no confounding.)
- A logistic regression without covariates gives the *marginal* odds ratio.
- However, if you add outcome predictors to this model, **the parameter of interest changes**.
  - This means that what you mean by « effect » has changed, even though the true effect does not change.
  - This also means that two analysts may choose to adjust for different causes of the outcome and therefore target two different parameter values.

# Non-collapsibility without confounding

- Consider a situation where there are no confounders of the  $A$ - $Y$  association.
  - For instance, in an clinical trial where  $A$  is randomized.
  - $Y$  is a binary outcome
- $C$  is a pure cause of the outcome.



# Comparison between conditional and marginal effect

- We have that  $Y$  is generated conditional on  $A$  and  $C$ :

$$\text{logit}P(Y = 1|A, C) = b + b_1A + b_2C$$

–  $\exp(b_1)$  is the *conditional odds ratio*

- The *marginal odds ratio*  $\exp(b_1^*)$  is defined as:

$$\text{logit}P(Y = 1|A) = b^* + b_1^*A$$

$$\exp(b_1^*) = \frac{P(Y^1 = 1)/P(Y^1 = 0)}{P(Y^0 = 1)/P(Y^0 = 0)}$$

# Comparison between conditional and marginal effect

- Goal: to compare the *true* values of  $\exp(b_1)$  and  $\exp(b_1^*)$ .
  - This will demonstrate the difference in the *targeted* parameters.
  - With a very large sample size, this represents the estimation bias we would obtain using a conditional logistic regression when the goal is really to estimate the marginal odds ratio.
- We can vary  $b$ ,  $b_1$ , and  $b_2$

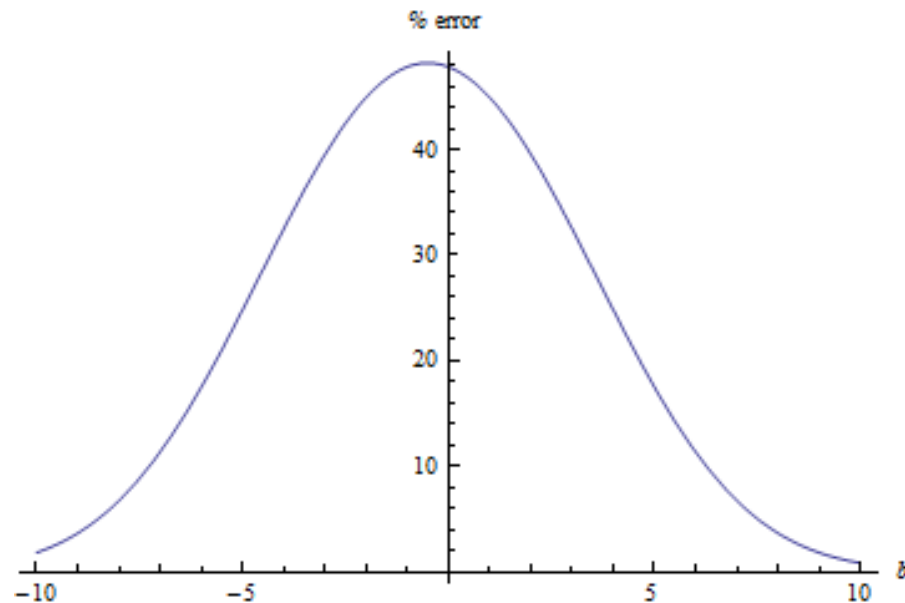
# Relative bias measure

- Percent error of the conditional parameter

$$\frac{\exp(b_1) - \exp(b_1^*)}{\exp(b_1^*)} * 100$$

# Varying the baseline risk

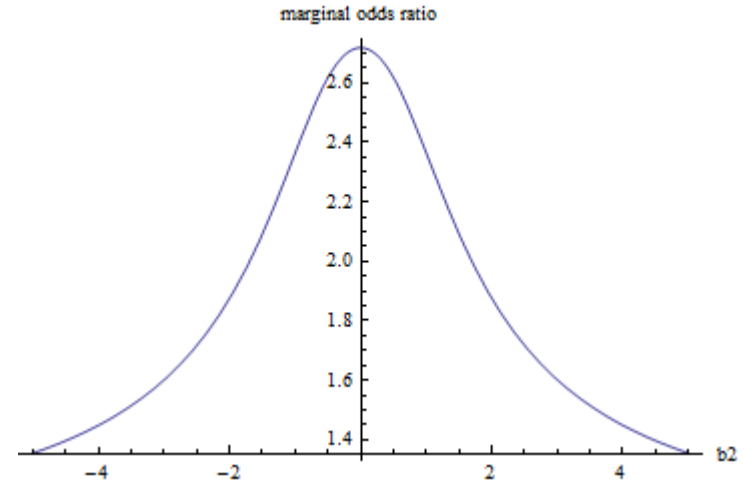
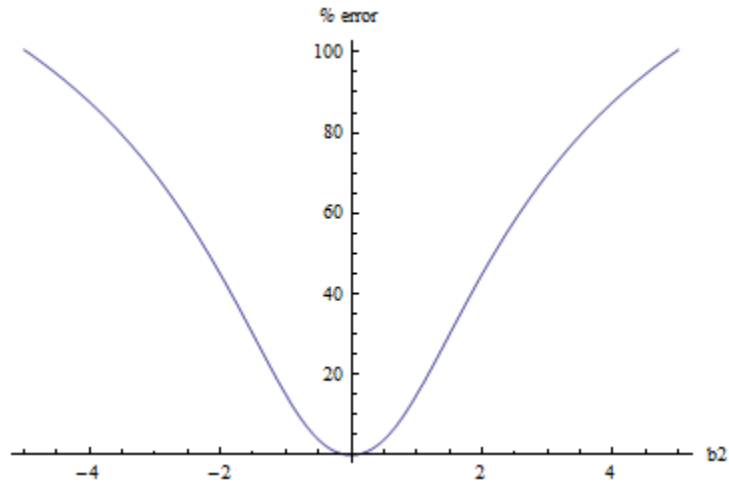
- Let  $b_1 = 1$  and  $b_2 = 2$ .



- As the outcome gets extremely rare or common,  $OR \approx RR$ .
- In the presence of a risk predictor with  $OR = \exp(2)$ , bias disappears when the baseline *risk* approaches  $\exp(-10) = 0.000045$

# Varying the strength of the risk predictor

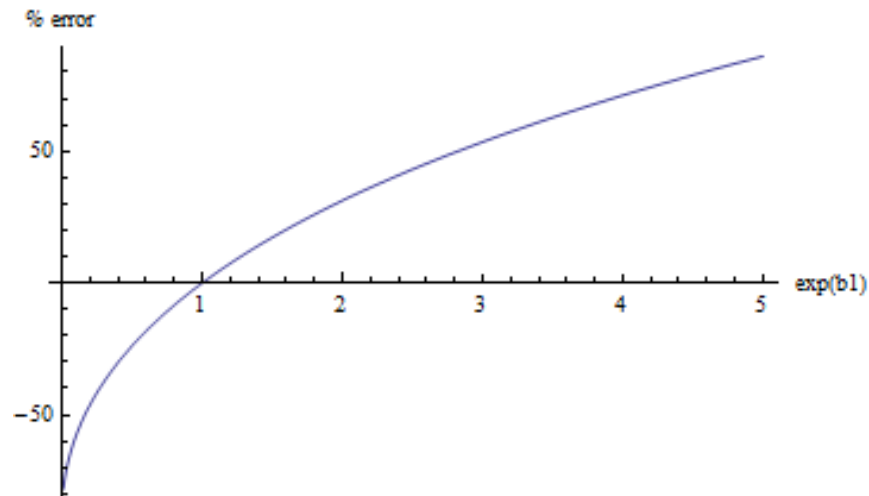
- Let  $b_1 = 1$  and  $b = 0$



- The marginal odds ratio decreases when the strength of the risk factor increases.
- The % error increases as the strength of the risk predictor increases (no bias when it is not a predictor).

# Varying the conditional odds ratio

- Let  $b = 0$  and  $b_2 = 2$

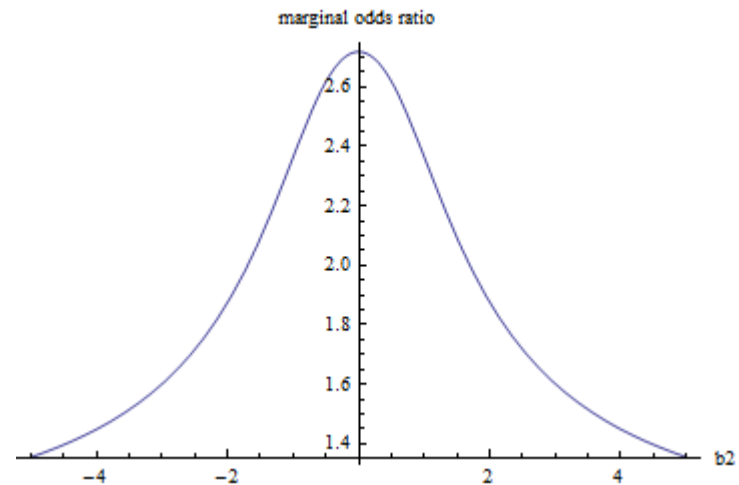
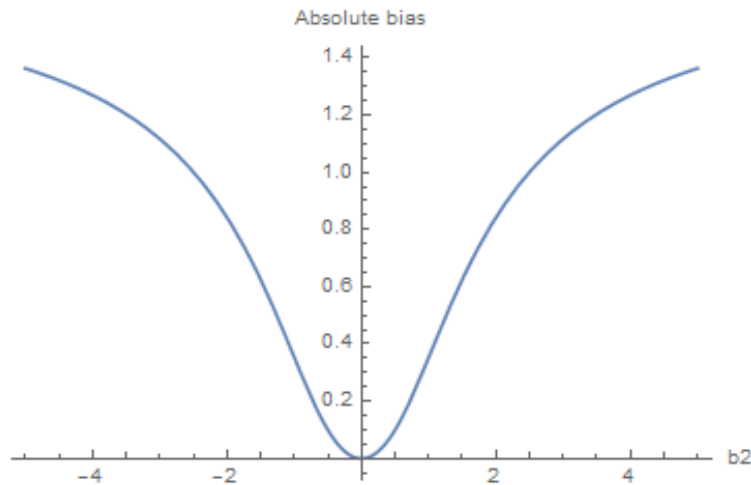


- For negative effects, the bias is *negative*.
- For positive effects, the bias is *positive*.
- Regardless of direction of risk predictor, the bias is always *anti-conservative*.



# Absolute bias when varying the strength predictor

- To see the magnitude of the bias, define the absolute bias as  $\exp(b_1) - \exp(b_1^*)$ .
- Let  $b_1 = 1$  and  $b = 0$



# Additional issues

- Because using a logistic regression for the estimation of the conditional odds ratio assumes a full parametric structure, confidence intervals will be narrow.
  - In general parametric assumptions are likely flawed.
  - So not only will the conditional odds ratio always inflate the apparent magnitude of the association, but it will also be « overly confident » about it.
- In the numerical examples, we assumed that the conditional logistic regression was correct.
  - In practice, it is not, potentially leading to additional bias.

# But what if I want the conditional OR?

- If you are interested in effect modification, you might reasonably fit

$$\text{logit}P(Y = 1|A, C) = b + b_1A + b_2C + b_3AC$$

in order to investigate  $b_3$ .

- However, I argue that all real-world interpretations of  $b_1$  and  $b_1^*$  are indifferntiable.
- I.e. if you have only main-terms in the model, you are essentially interpreting  $b_1$  as  $b_1^*$ . And this can be entirely misleading.

# Do your own investigation

- I provide the Mathematica code that I used to compute the true values of the marginal and conditional odds ratios and make the plots.
- Change the parameter values (of  $b$ ,  $b_1$ , and  $b_2$ ) and see how this affects the results.
- You can also investigate this issue using simulated datasets with large sample sizes.